

Discussion
Section,
week of
11-15 May
2020

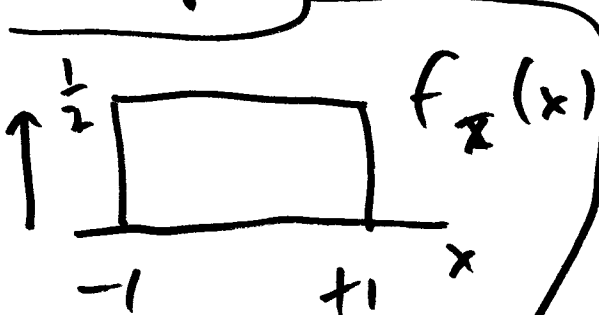
Transformations of random variables

STAT 131
11 May 20
DD disc.
section

You're working with a continuous ^①
random variable X with PDF

$f_X(x)$, and you get interested in a new
random variable Z that's related to
 X via $Z \equiv g(X)$ for some invertible
function $g(\cdot)$ - Z is called a transformed
version of X ; how does the probability
behavior of Z relate to that of X ?

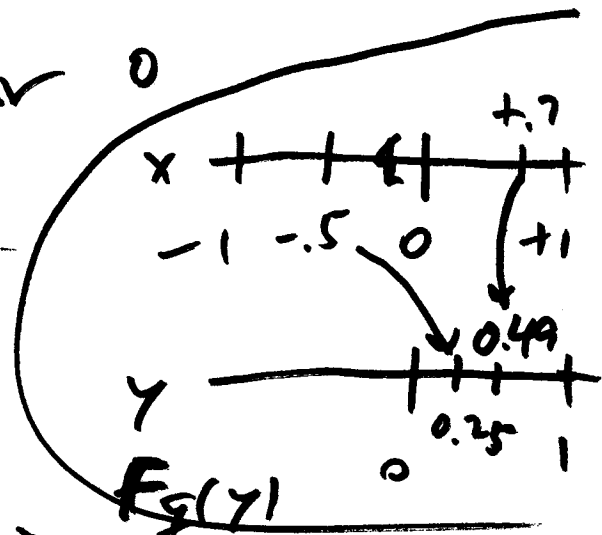
Example $X \sim \text{Uniform}([-1, +1])$ (continuous)
Support of X is S'_X



Set $Z = X^2$; PDF of Z !

First thing to note: $f'_Z = [0, 1]$ (2)
 ↗ support of Z

Intuition: when you square an x between -1 and $+1$, the result moves closer to 0 , so the PDF of Z should have lots of probability near 0



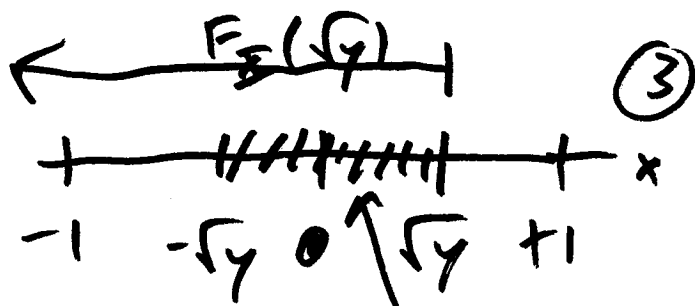
Method to get PDF of

$Z = g(X)$, first get CDF of Z in terms of CDF of X and then (2) differentiate $F_Z(y)$ in y to get PDF $f_Z(y)$ (for any $0 < y < 1$)

$$F_Z(y) = P(Z \leq y) = P(X^2 \leq y)$$

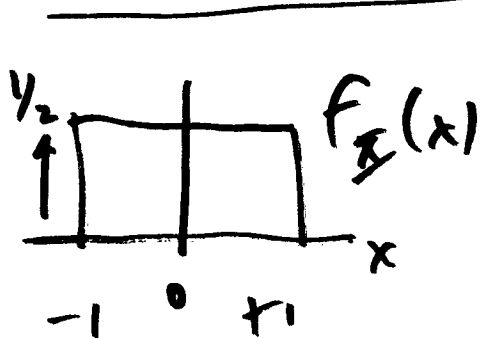
$$= P(-\sqrt{y} \leq Z \leq \sqrt{y})$$

(inverting $g(Z)$). $\textcircled{*}$



$$= F_Z(\sqrt{y}) - F_Z(-\sqrt{y})$$

this is the probability we want $\textcircled{*}$



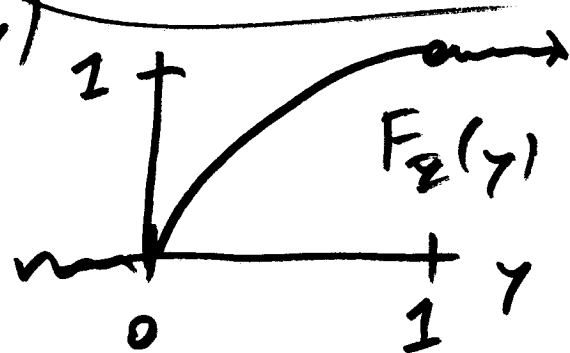
$$F_Z(x) = \begin{cases} 0 & \text{for } x \leq -1 \\ \int_{-1}^x \frac{1}{2} dt = \frac{x+1}{2} & \text{for } -1 \leq x \leq 1 \\ 1 & \text{for } x \geq 1 \end{cases}$$

for $(0 \leq y \leq 1)$

$$\text{So } F_Z(y) = F_Z(\sqrt{y})$$

$$- F_Z(-\sqrt{y})$$

$$= \left(\frac{1+\sqrt{y}}{2}\right) - \left(\frac{1-\sqrt{y}}{2}\right) = \sqrt{y}$$

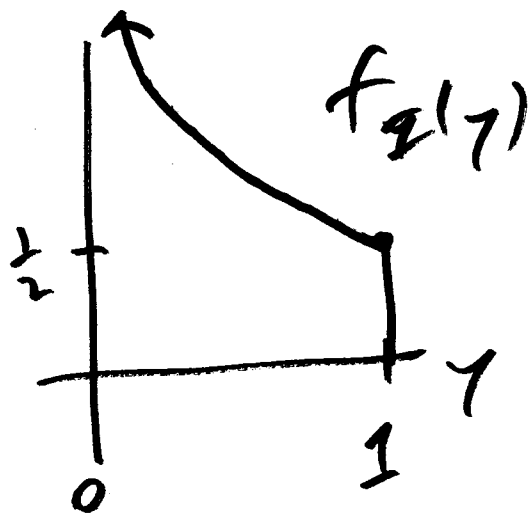


we could also have computed this as

$$P(-\sqrt{y} \leq Z \leq \sqrt{y}) = \int_{-\sqrt{y}}^{+\sqrt{y}} \frac{1}{2} dt = \sqrt{y} \checkmark$$

so for $0 \leq y \leq 1$, $F_{\mathcal{I}}(y) = \sqrt{y}$ and (4)

therefore $f_{\mathcal{I}}(y) = \frac{d}{dy} \sqrt{y} = \frac{d}{dy} y^{\frac{1}{2}} = \frac{1}{2\sqrt{y}}$



Remarkably, the PDF of $\mathcal{I} = \mathcal{X}^2$ is unbounded at 0: $\lim_{y \downarrow 0} \frac{1}{2\sqrt{y}} = +\infty$ (!!)

(R demo)

And now, for an amazing use of this idea:

general setup: \mathcal{X} continuous r.v. with CDF $F_{\mathcal{X}}(x)$ and PDF $f_{\mathcal{X}}(x)$; $\mathcal{I} \doteq g(\mathcal{X})$ for

invertible $g(\cdot)$ My favorite invertible

function in this class is $g(x) = F_{\mathcal{X}}(x)$; what happens with that choice of $g(\cdot)$?

$Z = F_X(X)$ (!) | Let's just see ^{what} happens: (5)

note: $S_X = (0, 1)$

$$F_Z(y) = P(Z \leq y) = P[F_X(X) \leq y]$$

Now, since X is continuous, $F_X(\cdot)$ and strictly increasing is invertible on the support S_X of X ,

$$\text{so } F_Z(y) = P[F_X(X) \leq y] = P[X \leq F_X^{-1}(y)]$$

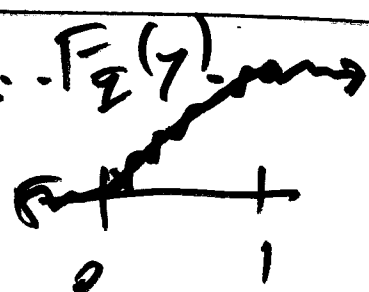
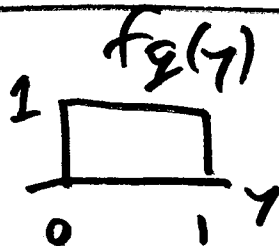
$$\stackrel{(!)}{=} F_X[F_X^{-1}(y)] \stackrel{(!)}{=} y$$

But there's only

one random variable Z with CDF

$$F_Z(y) = y \text{ (for } 0 < y < 1), \text{ namely}$$

Uniform $(0, 1)$ (!)



So what we have just shown is ⑥

that if X has CDF $F_X(\cdot)$ then

$$Z = F_X(X) \sim \text{Uniform}(0, 1) (!)$$

Now, to finish off the argument by creating something extremely useful in data science:

let's run this argument backwards

$$Z = F_X(X)$$

↑

Uniform(0, 1) transformed random variable

by hitting a Uniform(0, 1) $F_X^{-1}(Z)$; random variable with $F_X^{-1}(\cdot)$,

do we recover X ?

let's see:

let $U \sim \text{Uniform}(0, 1)$ and create ⑦

$W = F_X^{-1}(U)$; what's the CDF of W ?

$$F_W(w) = P(W \leq w) = P[F_X^{-1}(U) \leq w]$$

But in this setup, because X is continuous,

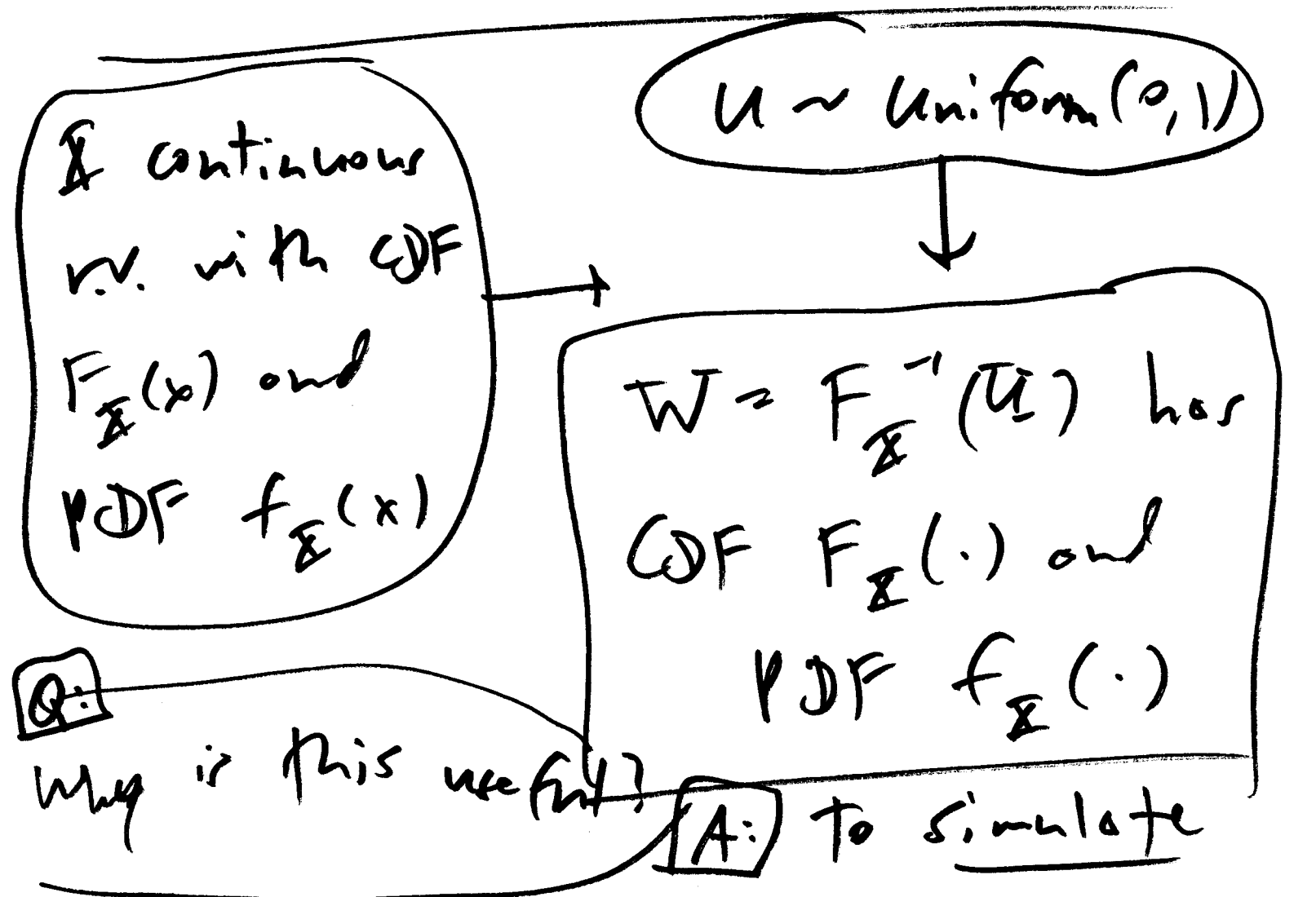
$F_X^{-1}(\cdot)$ is also invertible and

strictly increasing. So we get

$$F_W(w) = P[F_X^{-1}(U) \leq w] = P[U \leq F_X(w)]$$

But $0 \leq F_X(w) \leq 1$, and we already worked out that a $\text{Uniform}(0, 1)$ rv such as U has CDF $P(U \leq u) = u$ (for $0 \leq u < 1$)

And finally, then, $F_W(w) = F_Z(z)$; $\textcircled{5}$
(continuous)
two rv with the same CDF must have
the same PDF; so (the envelope, please)



random draws (the foundation of
the Monte Carlo method for learning
about the world) from a PDF $f_Z(x)$.

Algorithm (Inverse Probability Integral Transform) ⁽⁹⁾

① Given $f_X(x)$, compute $F_X(x)$ and set $F_X(x) = u$

② Compute $x_p = F_X^{-1}(p)$ by solving for x_p

③ generate $U_1^*, \dots, U_n^* \sim \text{Uniform}(0, 1)$

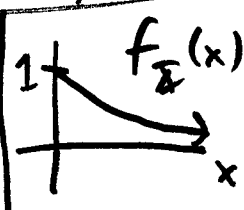
④ Set $X_i^* = F_X^{-1}(U_i^*)$

Example

(extra notes p. 81)

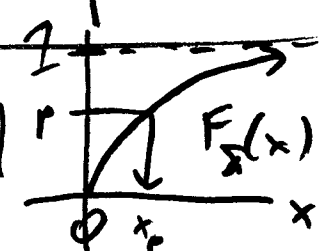
X = voltage in an electrical system

$$f_X(x) = \begin{cases} \frac{1}{(1+x)^2} & \text{for } x > 0 \\ 0 & \text{else} \end{cases}$$



① for $x > 0$

$$F_X(x) = \int_0^x \frac{1}{(1+t)^2} dt = \frac{x}{x+1} \quad \boxed{Wd}$$



② $F_{\mathcal{I}}(x_p) = p \rightarrow \frac{x}{p} = p \xrightarrow{\text{solve for } x_p} \boxed{x_p = \frac{p}{1-p}} \textcircled{10}$

(odds ratio (1))

③ Generate $U_i^* \stackrel{\text{i.i.d.}}{\sim} \text{Uniform}(0, 1)$
for $i = 1, \dots, M$

(R demo)

④ Compute $\mathcal{I}_i^* = \frac{U_i^*}{1 - U_i^*} \quad (i = 1, \dots, M)$
(R demo)

In the early days of computing (mid to late 1940s for the theory (Turing, von Neumann, ...), early 1950s for the practice, pseudo-uniform "random" number generators were the first Monte Carlo algorithms created; IPIT greatly expanded the scope of simulation.