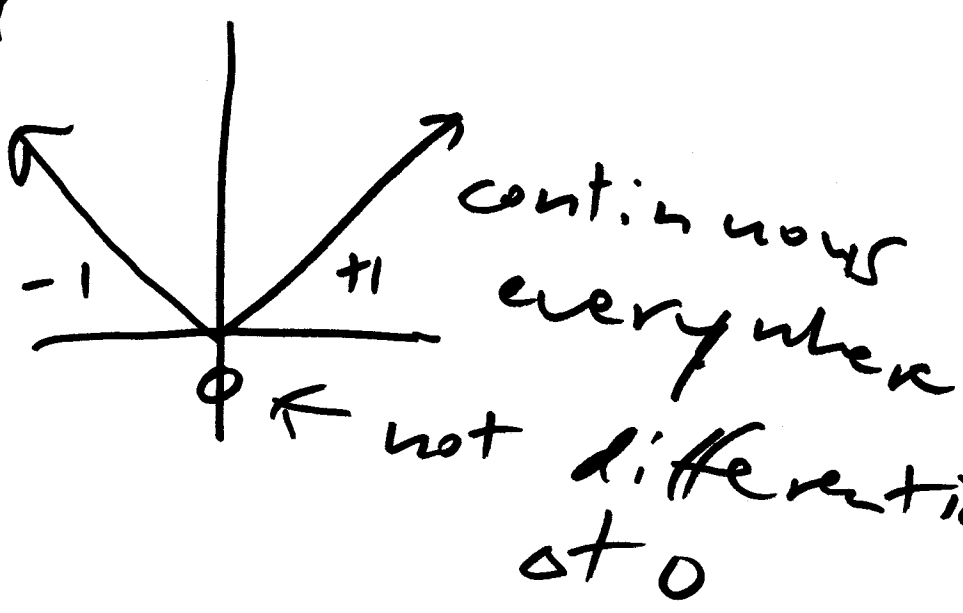


STAT 131  
21 May 20

(D) office  
1.5 hour

TAT 2#3 (1)



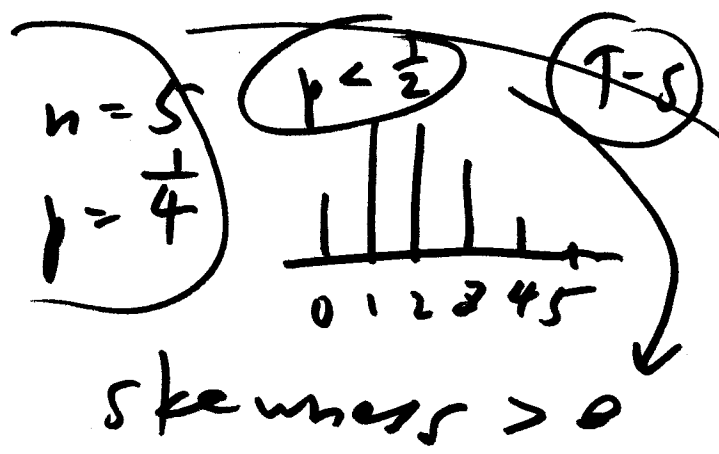
$X \sim$   
Binomial( $n, p$ )

$0 < p < 1$   
 $n \geq 3$

$E(X) = np = \mu$   
 $E(X^2) = np[1 + (n-1)p]$   
 $E(X^3) =$

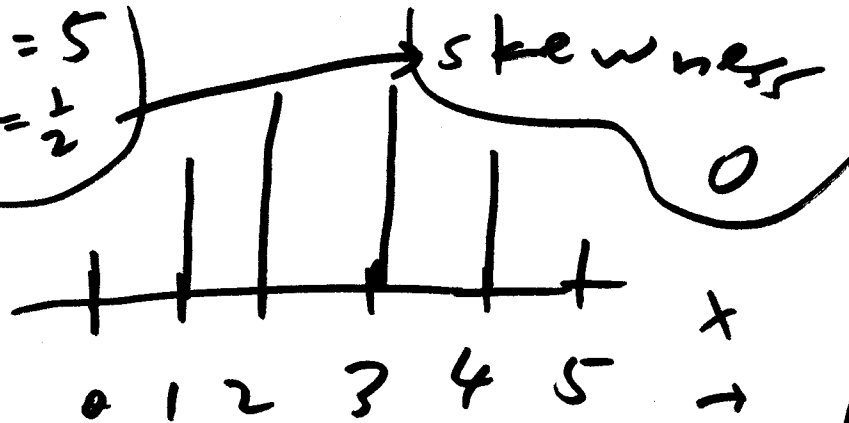
$np[1 + (n-1)(n-2)p^2 + 3p(n-1)]$

$V(X) = np(1-p) = \sigma_X^2$   
 $X$  has P.M.F



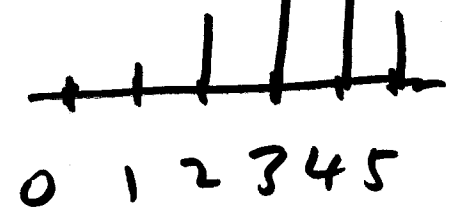
$f_X(x | n, p) =$   
 $\begin{cases} \binom{n}{x} p^x (1-p)^{n-x} & \text{for } x=0, 1, \dots, n \\ 0 & \text{else} \end{cases}$

$n=5$   
 $p=\frac{1}{2}$



Symmetric  
about  $x=2.5$

$(n=5, p=\frac{3}{4})$  (2)  
(skewness < 0)



mirror image  
of  $(n=5, p=\frac{1}{4})$   
reflected about  
 $x=2.5$

$$\text{skewness}(\mathcal{X}) = E\left[\left(\frac{\mathcal{X} - \mu_{\mathcal{X}}}{\sigma_{\mathcal{X}}}\right)^3\right]$$

$$\begin{aligned} \sigma(\mathcal{X}) &= \sigma_{\mathcal{X}} \\ &= \sqrt{V(\mathcal{X})} \\ &= \sqrt{np(1-p)} \end{aligned}$$

$$\begin{aligned} &= \frac{1}{\sigma_{\mathcal{X}}^3} E\left[(\mathcal{X} - \mu_{\mathcal{X}})^3\right] \\ &= \frac{1}{\sigma_{\mathcal{X}}^3} E\left(\mathcal{X}^3 - 3\mathcal{X}^2\mu_{\mathcal{X}} + 3\mathcal{X}\mu_{\mathcal{X}}^2 - \mu_{\mathcal{X}}^3\right) \end{aligned}$$

$$\text{skewness}(X) = \frac{1}{\sigma_X^3} \left\{ E(X^3) + E(\underline{-3\mu_X X^2}) \right. \\ \left. + E(\underline{3\mu_X^2 X}) + E(\underline{-\mu_X^3}) \right\} \quad (3)$$

$$= \frac{1}{\sigma_X^3} \left\{ E(X^3) - 3\mu_X E(X^2) \right. \\ \left. + 3\mu_X^2 E(X) - \mu_X^3 \right\}$$

(hideous algebra)

=

$$\frac{1 - 2p}{\sqrt{4p(1-p)}} = 0 \iff$$

$$1 - 2p = 0$$

$$\iff p = \frac{1}{2}$$

plugging  
in exp. for

$E(X), \dots$

proof by wa

all Binomial distributions  
with  $p = \frac{1}{2}$  are symmetric,  
no matter what  $n$  is

$$\text{skewness}(\mathbb{X}) = \frac{1-2p}{\sqrt{4p(1-p)}} \rightarrow 0 \quad (4)$$

as  $n \rightarrow \infty$  for all  $0 < p < 1$

skewness  $(\mathbb{X}) \rightarrow 0$  as  $n \rightarrow \infty$

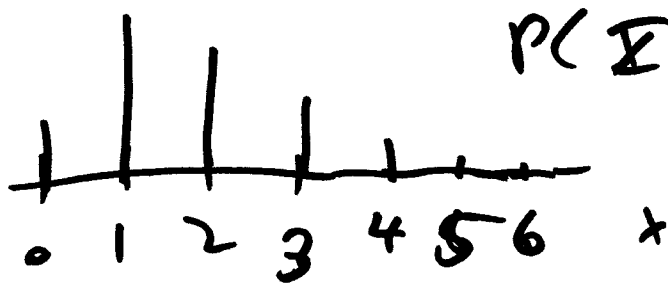
discrete

$$(\mathbb{X} | \lambda) \sim \text{Poisson}(\lambda) \leftrightarrow$$

$$(\lambda > 0) \quad \begin{matrix} \uparrow \\ f_{\mathbb{X}}(x | \lambda) = \begin{cases} \frac{\lambda^x e^{-\lambda}}{x!} & \text{for } x = 0, 1, \dots \\ 0 & \text{else} \end{cases} \\ \uparrow \\ \text{PMF} \\ P(\mathbb{X} = x | \lambda) \end{matrix}$$

$$S_{\mathbb{X}} = \{0, 1, \dots\}$$

Poisson can model  $\mathbb{X} = \left( \begin{matrix} \# \text{ of interesting} \\ \text{occurrences} \\ \text{in } (0, T) \end{matrix} \right)$



$$P(X=x | \lambda=1.1)$$

PMF

MGF (5)  
MGF?

$X$  discrete

$$\psi_X(x) = E(\underline{e^{tx}})$$

(LOTUS)

$$= \sum_{x=0}^{\infty} e^{tx} \cdot P(X=x | \lambda)$$

$$= \sum_{x=0}^{\infty} e^{tx} \cdot \frac{\lambda^x e^{-\lambda}}{x!}$$

$$= e^{-\lambda} \sum_{x=0}^{\infty} \frac{e^{tx} \cdot \lambda^x}{x!}$$

$$= e^{-\lambda} \sum_{x=0}^{\infty} \frac{(\lambda e^t)^x}{x!}$$

$$= e^{-\lambda} \exp(\lambda e^t)$$

$$= e^{-\lambda} e^{\lambda e^t} = e^{-\lambda + \lambda e^t}$$

$$= e^{\lambda(e^t - 1)}$$

$$= e^{\lambda(e^t - 1)} \quad \checkmark$$

for any  $w \in \mathbb{R}$

$$e^w = \sum_{x=0}^{\infty} \frac{w^x}{x!}$$